

$Y = u(X)$, $u(x)$ diff & either inc. or dec. $\Rightarrow u'$ exists.

$$\text{then } f_Y(y) = f_X(u^{-1}(y)) \left| \frac{d}{dy}(u^{-1}(y)) \right|$$

Ex: let $Z \sim N(0,1) \Rightarrow f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$

$$\text{Define } Y = Z^2 \Rightarrow Z = \sqrt{Y} \text{ if } Z > 0. \\ = -\sqrt{Y} \text{ if } Z < 0.$$

$$\frac{d}{dy}(\sqrt{y}) = \frac{1}{2\sqrt{y}}, \quad \frac{d}{dy}(-\sqrt{y}) = -\frac{1}{2\sqrt{y}}$$

$$f_Y(y) = f_Z(\sqrt{y}) \left| \frac{1}{2\sqrt{y}} \right| + f_Z(-\sqrt{y}) \left| \frac{-1}{2\sqrt{y}} \right|$$

$$= \frac{1}{\sqrt{2\pi}} e^{-y/2} \frac{1}{2\sqrt{y}} \cdot 2$$

$$= \frac{e^{-y/2}}{\sqrt{2\pi y}} \text{ for } y > 0$$

$$= \frac{1}{\sqrt{2}\sqrt{\pi}} y^{-1/2} e^{-y/2}$$

$$= \frac{1}{2^{1/2}\sqrt{\pi}} y^{\frac{1}{2}-1} e^{-\frac{y}{2}}$$

$$= \frac{1}{\Gamma(\frac{1}{2}) 2^{1/2}} y^{\frac{1}{2}-1} e^{-\frac{y}{2}}$$

so $Y \sim \text{Gamma}(\frac{1}{2}, 2)$

$$Y \sim \chi_1^2$$

7.4 Transformation Tech. Multiple Vars

Let $X \sim P_0(\theta)$, $Y \sim P_0(\lambda)$, X, Y independent.

$$f_{xy}(x,y) = \left(\frac{e^{-\theta} \theta^x}{x!} \right) \left(\frac{e^{-\lambda} \lambda^y}{y!} \right) \quad \text{for } x=0, 1, 2, \dots \\ y=0, 1, 2, \dots$$

Let $U = X+Y$, $V = Y \rightarrow$ if $Y=12$ and want $P(U=25)$, this is $P(Y=12)$

$$\text{so } Y=V, \quad X=U-Y=U-V.$$

$$V=0, 1, 2, \dots$$

$$U = V, V+1, V+2, \dots$$

$$f_{uv}(u,v) = f_{xy}(u-v, v) = \frac{e^{-\theta} \theta^{u-v}}{(u-v)!} \frac{e^{-\lambda} \lambda^v}{v!}$$

$$\begin{aligned} f_u(u) &= \sum_{v=0}^u f_{xy}(u-v, v) = \sum_{v=0}^u \frac{e^{-\theta-\lambda} \theta^{u-v} \lambda^v}{(u-v)! v!} \\ &= \frac{e^{-(\theta+\lambda)}}{u!} \sum_{v=0}^u \frac{u!}{(u-v)! v!} \lambda^v \theta^{u-v} \\ &= \frac{e^{-(\theta+\lambda)} (\theta+\lambda)^u}{u!} \quad \text{binom. thm.} \end{aligned}$$

$$\text{So } U = P_0(\theta + \lambda)$$