

Lec 10/17

Monday, October 17, 2016 8:02 AM

$$X = X_1 + \dots + X_n$$

where $X_i \stackrel{i.i.d.}{\sim} \text{Bern}(\theta)$

$$\Rightarrow X \sim \text{Bin}(n, \theta)$$

$$P(X) = \begin{cases} \binom{n}{x} \theta^x (1-\theta)^{n-x} & \text{if } x = 0, \dots, n \\ 0 & \text{o.w.} \end{cases}$$

$$E(X) = n\theta \quad \text{Var}(X) = n\theta(1-\theta)$$

5.5 Geometric & negative Binomial dists.

Consider a scenario where you conduct a sequence of ^{independent} Bernoulli trials until you see 1st success.

E.g.: shoot free throws until first make.

$$P(X) = \begin{cases} \theta(1-\theta)^{x-1} & \text{if } x = 1, 2, \dots \\ 0 & \text{o.w.} \end{cases}$$

$$X \sim \text{Geom}(\theta)$$

$$\begin{array}{ccc} \theta & (1-\theta)^{(x-1)} & \\ \downarrow & \swarrow & \\ \text{1 success} & & x-1 \text{ failures} \end{array}$$

$$M_X(t) = E(e^{tx}) = \sum_{x=1}^{\infty} e^{tx} (1-\theta)^{x-1} \theta = \theta e^t \sum_{x=1}^{\infty} (e^t(1-\theta))^{x-1}$$

$$= \theta e^t \sum_{x=1}^{\infty} (e^t(1-\theta))^x = \theta e^t \left(\frac{1}{1 - (e^t(1-\theta))} \right) \quad \text{if } e^t(1-\theta) < 1$$

$$= \frac{\theta e^t}{1 - (1-\theta)e^t}$$

$$e^t < \frac{1}{1-\theta}$$

$$t < -\log(1-\theta)$$

$\psi(t) = \log(M_X(t))$ (cumulative generating function)

$$\psi'(0) = \mu \quad \psi''(0) = \sigma^2$$

$$E(X) = \frac{1}{\theta}$$

$$\text{Var}(X) = \frac{1-\theta}{\theta^2}$$

Negative Binomial

$X = \#$ trials required to get k^{th} success:

$$P(X) = \begin{cases} \binom{x-1}{k-1} \theta^k (1-\theta)^{x-k} & \text{for } x = k, k+1, \dots \\ 0 & \text{o.w.} \end{cases}$$

$$X \sim \text{NegBin}(k, \theta)$$

Note: Geometric is negative binomial where $k=1$.

$$M_X(t) = E(e^{tx}) = \sum_{x=k}^{\infty} e^{tx} \binom{x-1}{k-1} \theta^k (1-\theta)^{x-k}$$

$$(e^t)^x = (e^t)^{x-k+k}$$

$$= (e^t)^{x-k} e^{tk}$$

$$M_X(t) = E(e^{tx}) = \sum_{x=k}^{\infty} e^{tx} \binom{x-1}{k-1} \theta^k (1-\theta)^{x-k}$$

$$= (e^t)^{x-k} e^{tk}$$

$$= \sum_{x=k}^{\infty} \binom{x-1}{k-1} (\theta e^t)^k (1-\theta e^t)^{x-k}$$

$$= \frac{(\theta e^t)^k}{(1-(1-\theta)e^t)^k} \sum_{x=k}^{\infty} \binom{x-1}{k-1} [1-(1-\theta)e^t]^k [(1-\theta)e^t]^{x-k} \rightarrow \text{this is sum of pmf values, so it is 1.}$$

$\downarrow \theta'$ $\downarrow 1-\theta'$

$$= \left[\frac{\theta e^t}{1-(1-\theta)e^t} \right]^k$$

θ' must be in $[0, 1]$

$$0 \leq (1-\theta)e^t \leq 1$$

$$0 \leq e^t \leq \frac{1}{1-\theta}$$

$$t \leq -\log(1-\theta)$$

So $E(x) = \frac{k}{\theta}$

$$\text{Var}(x) = \frac{k(1-\theta)}{\theta}$$

Ex: Suppose 6% of the U.S. pop has A⁻ blood type.

Assume independence from donor to donor.

a) Probability the first A⁻ donor is the 12th

b) Probability the second A⁻ donor is the 20th

a) $X = \# \text{ donors until 1st A}^- \text{ donor.}$

$$P(X=12) = 0.06 \cdot 0.94^{11} = 0.0303...$$

b) Let $Y = \# \text{ donors until 2nd A}^- \text{ donor}$

$$Y \sim \text{Neg Bin}(2, 0.06)$$

$$P(Y=20) = \binom{14}{1} (0.06)^2 (0.94)^{12}$$

$$= 0.0225$$